Parameterized Property Testing of Functions

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Abstract

We investigate the parameters in terms of which the complexity of sublinear-time algorithms should be expressed. Our goal is to find input parameters that are tailored to the combinatorics of the specific problem being studied and design algorithms that run faster when these parameters are small. This direction enables us to surpass the (worst-case) lower bounds, expressed in terms of the input size, for several problems. Our aim is to develop a similar level of understanding of the complexity of sublinear-time algorithms to the one that was enabled by research in parameterized complexity for classical algorithms.

Specifically, we focus on testing properties of functions. By parameterizing the query complexity in terms of the size $r$ of the image of the input function, we obtain testers for monotonicity and convexity of functions of the form $f : [n] \rightarrow \mathbb{R}$ with query complexity $O(\log r)$, with no dependence on $n$. The result for monotonicity circumvents the $\Omega(\log n)$ lower bound by Fischer (Inf. Comput., 2004) for this problem. We present several other parameterized testers, providing compelling evidence that expressing the query complexity of property testers in terms of the input size is not always the best choice.

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1 Introduction

In this paper, we set out to investigate the parameters in terms of which the complexity of sublinear-time algorithms should be expressed. Our goal is to find input parameters that are tailored to the combinatorics of the specific problem being studied and design algorithms that run faster when these parameters are small. This direction could enable one to surpass the (worst-case) lower bounds on the problem complexity that are usually expressed in terms of the input size. The spirit of our study is similar to that in the field of parameterized complexity. In parameterized complexity, the focus is on expressing the complexity of problems as a function of one or more input parameters in order to obtain a fine-grained complexity classification, for example, of NP-hard problems. Our aim is to develop a similar level of understanding of the complexity of sublinear-time algorithms to the one that was enabled by research in parameterized complexity for classical algorithms.

We focus our study on the framework of property testing, introduced by Goldreich et al. [28] and Rubinfeld and Sudan [40]. In property testing, an algorithm (an $\varepsilon$-tester) for
property $\mathcal{P}$, where $\mathcal{P}$ is viewed as a class of functions, is given a parameter $\varepsilon \in (0,1)$ as input and has oracle access to a function $f$. The tester has to accept with probability at least $2/3$ if $f$ belongs to the class $\mathcal{P}$, and reject with probability at least $2/3$ if $f$ is $\varepsilon$-far from $\mathcal{P}$, that is, differs from every function in $\mathcal{P}$ on at least an $\varepsilon$ fraction of function values. In the context of property testing of functions, the query complexity of a tester is usually expressed in terms of $\varepsilon$ and the size of the domain of the input function. This works well for properties whose query complexity depends only on the proximity parameter $\varepsilon$. However, for other properties, it is not clear whether the domain size is the right parameter to express their testing complexity.

Consider, for example, the widely studied problem of testing monotonicity of real-valued functions (see, e.g., [27, 22, 23, 36, 26, 24, 31, 1, 32, 2, 11, 10, 13, 16, 12, 9, 17, 18, 15, 20, 19, 35, 4, 5, 21], and recent surveys [38, 14]). For functions over a discrete domain $[n]$ (also called the line), monotonicity testing is equivalent to testing sortedness of arrays. Algorithms for sortedness testing have found use, for instance, in determining the “state of sortedness” of relational databases [6], where the testing step is performed to decide on the sorting algorithms to be run on the database. The complexity of sortedness testing (for constant $\varepsilon$) is $\Theta(\sqrt{n})$ if the tester is only allowed to make independent and uniformly random queries [26]; it is $\Theta(\log n)$ if the tester is allowed to make arbitrary queries [23, 24].

From the above discussion, it might appear that one cannot make any more improvements to the complexity of monotonicity testing over $[n]$. However, we argue that this is the case only when the complexity of the problem is parameterized in terms of $n$, the domain size.

In this work, we ask whether better monotonicity testers can be designed by parameterizing the query complexity in terms of the size of the image of the input function. The starting point for our investigation is the folklore result that, for $\varepsilon$-testing monotonicity of Boolean functions over $[n]$, only $O(1/\varepsilon)$ queries suffice. A slightly more general corollary of this result is that monotonicity of functions over $[n]$ with image size at most $2$ can be $\varepsilon$-tested with only $O(1/\varepsilon)$ queries. The only bound for monotonicity testing (over $[n]$) that is expressed in terms of the image size $r$ of the input function is the bound of $\Omega(\log r)$ for nonadaptive\footnote{Testers whose queries do not depend on the answers to previous queries are called \emph{nonadaptive}; general testers that do not satisfy this requirement are \emph{adaptive}.} testers due to Blais et al. [12]. We design an $\varepsilon$-tester for monotonicity of functions over $[n]$ with query complexity $O((\log r)/\varepsilon)$, where $r$ is an upper bound on the size of the image of the input function. This result circumvents Fischer’s lower bound of $\Omega(\log n)$ for this problem by focusing on a different parameter for measuring query complexity.

The size of the image is one of the natural parameters in terms of which one can express the complexity of property testing algorithms. In this work, we show that there are several testing problems for which parameterizing the complexity in terms of the image size works well. Another example where parameterization has helped in the design of efficient testers is the work of Jha and Raskhodnikova [34] on Lipschitz testing, even though they do not view their results from this angle. The complexity of their testers is expressed in terms of the \emph{image diameter}. The image diameter of a function $f: \mathcal{D} \to \mathbb{R}$ is $\max_{x, y \in \mathcal{D}} |f(x) - f(y)|$. In many situations, the image diameter is much smaller than the domain size. We believe that all this evidence is compelling enough to make one rethink the way in which the complexity of sublinear-time algorithms is expressed. Our paper is a first step towards formalizing this notion and finding what we think are the right parameters to express the complexity of some central problems in sublinear-time algorithms.
1.1 Parameters and Properties Studied in this Work

We study the dependence of complexity of monotonicity and convexity testers on the image size of the input functions. The image of a function is defined as follows.

Definition 1.1 (Image of a function). Let $f$ be a function defined over a finite, discrete domain $D$. The image of $f$, denoted $\text{Im}(f)$, is the set $\{f(x) : x \in D\}$ or, in other words, the set of all values taken by $f$ on points in $D$.

For the special case, when $D$ is $[n]$, a function $f : [n] \rightarrow \mathbb{R}$ can also be viewed as a real-valued array of length $n$. Here, $\text{Im}(f)$ is equal to the set of distinct values in the array.

We restrict our attention to real-valued functions defined over the following domains. These are domains for which testing monotonicity and convexity have been studied extensively.

Definition 1.2 (Hypergrid, Hypercube, Line). For $x \in [n]^d$, let $x_i$ denote the $i^{th}$ coordinate of $x$. A hypergrid is a partial order $([n]^d, \preceq)$ where $x \preceq y$ means that $x_i \leq y_i$ for all $x, y \in [n]^d$ and $i \in [d]$. The partial order $([2]^d, \preceq)$ is called a hypercube and the total order $([n], \preceq)$ is called a line.

Next, we summarize some of the previous work on testing monotonicity and convexity of real-valued functions.

Monotonicity. A function $f : D \rightarrow \mathbb{R}$ defined over a partial order $(D, \preceq)$ is monotone if $f(x) \leq f(y)$ for all $x, y \in D$ satisfying $x \preceq y$. Monotonicity is one of the most widely studied properties in the field of property testing [27, 22, 23, 36, 26, 24, 31, 1, 32, 2, 11, 10, 13, 16, 12, 9, 17, 18, 15, 20, 19, 35, 4, 5, 21]. The complexity of $\varepsilon$-testing monotonicity of functions of the form $f : [n]^d \rightarrow \mathbb{R}$ is $\Theta\left(\frac{d \log n}{\varepsilon}\right)$ [16, 17]. For the special case of the line, the testing complexity is $\Theta\left(\frac{\log n}{\varepsilon}\right)$ [23, 24]. For functions defined over general poset domains $D$, the complexity of monotonicity testing is $O\left(\sqrt{|D|/\varepsilon}\right)$ [26].

Convexity. For a convex set $D$, a function $f : D \rightarrow \mathbb{R}$ is convex if $\forall x, y \in D$ and $t \in [0, 1]$, $f(tx + (1-t)y) \leq tf(x) + (1-t)f(y)$. For real-valued functions over $[n]$, convexity can be $\varepsilon$-tested using $O\left(\frac{\log n}{\varepsilon^2}\right)$ queries [37]. This bound is tight (for constant $\varepsilon$) for nonadaptive testers [12].

1.2 Our Results

In this section, we describe the key technical contributions of our work. We design efficient testers for monotonicity over various hypergrid domains and convexity over the line. For monotonicity of functions over the line $[n]$, which is equivalent to the property of sortedness of arrays of length $n$, we design efficient testers under two different models of input access: (i) query access and (ii) uniform samples. Our testers are given an upper bound $r$ on the image size of the input function, and their complexity is parameterized in terms of $r$.

Sortedness testing. We present our tester for sortedness of $n$-element arrays (monotonicity over the line $[n]$) in Section 3. The complexity of our tester is independent of $n$. Our tester has 1-sided error, that is, it always accepts a function with the property. (In contrast, the general tester is said to have 2-sided error.) We prove the following theorem.
Theorem 1.3. There exists a 1-sided error \( \varepsilon \)-tester making \( O \left( \frac{\log r}{\varepsilon} \right) \) queries to test sortedness of arrays with at most \( r \) distinct values.

An important ingredient in our sortedness tester is a nearly optimal nonadaptive tester for this task, presented in Section 2. Its performance is summarized in the next theorem.

Theorem 1.4. There exists a nonadaptive, 1-sided error \( \varepsilon \)-tester making \( O \left( \frac{1}{\varepsilon} \log \frac{r}{\varepsilon} \right) \) queries to test sortedness of arrays with at most \( r \) distinct values.

The query complexity of our nonadaptive tester matches (up to the dependence on \( \varepsilon \)) the \( \Omega(\log r) \) lower bound for nonadaptive sortedness testers in [12]. Note that for \( r \geq 1/\varepsilon \), the complexity of the nonadaptive tester is \( O \left( \frac{\log r}{\varepsilon} \right) \). The tester that we design to prove Theorem 1.3 runs the nonadaptive tester for \( r \geq 1/\varepsilon \) and a different (adaptive) tester, presented in Section 3, for \( r < 1/\varepsilon \).

Uniform sortedness testing. The work that defined property testing [28], in addition to the model with oracle access to the input, also considered testers that are allowed access to function values only at points sampled uniformly and independently at random from the domain. This model of property testing, known as uniform or sample-based testing, was further studied by Goldreich and Ron [30], Fischer et al. [25], Berman et al. [8] and Berman et al. [7]. The query complexity of \( \varepsilon \)-testing sortedness of \( n \)-element arrays (for constant \( \varepsilon \)) using only uniformly and independently drawn samples is \( \Theta(\sqrt{n}) \) [26]. We design optimal (up to the dependence on \( \varepsilon \)) uniform testers whose query complexity is parameterized in terms of the number or distinct elements in the input arrays. These results can be found in Sections 5 and 6.

Theorem 1.5. There exists a 1-sided error \( \varepsilon \)-tester that makes \( O(\sqrt{r}/\varepsilon) \) uniform and independent queries to test sortedness of arrays with at most \( r \) distinct values.

Theorem 1.6. Testing sortedness of arrays with values in \([r] \) requires \( \Omega(\sqrt{r}) \) uniform queries, even with 2-sided error.

Monotonicity testing over hypergrids. We present our tester for monotonicity of real-valued functions over hypergrid domains in Section 4 and prove the following theorem.

Theorem 1.7. There exists a 1-sided error \( \varepsilon \)-tester that makes \( O \left( \frac{d^2}{\varepsilon} \log \frac{d}{\varepsilon} \log r \right) \) queries to test monotonicity of real-valued functions \( f : [n]^d \mapsto \mathbb{R} \) over the hypergrid domain, where \( |\text{Im}(f)| \leq r \).

Note that our tester has a better complexity (up to log factors) than the optimal tester for monotonicity of real-valued functions over the hypergrid domains that makes \( O \left( \frac{d \log n}{\varepsilon} \right) \) queries [16] for small \( r \). Parameterizing the complexity of testing in terms of the image size of the functions being tested is what enables us to bypass the \( \Omega \left( \frac{d \log n}{\varepsilon} \right) \) lower bound for monotonicity testing of functions over hypergrid domains in [17].

Convexity testing over the line. Finally, in Section 7, we give a nonadaptive convexity tester for real-valued functions over the line and prove the following theorem.

Theorem 1.8. There exists a nonadaptive, 1-sided error \( \varepsilon \)-tester for convexity of functions \( f : [n] \mapsto \mathbb{R} \) that takes an integer \( r \geq |\text{Im}(f)| \) as input and makes \( O(1/\varepsilon) \) queries when \( r < \varepsilon n/3 \) and \( O \left( \frac{\log(1/\varepsilon)}{\varepsilon^2} \right) \) queries otherwise.
Recall that for real-valued functions over \([n]\), the complexity of (nonadaptively) \(\varepsilon\)-testing convexity (for constant \(\varepsilon\)) is \(\Theta(\log n)\). Contrary to this, our tester makes only a constant number of queries when the image size of the function is small.

### 1.3 Related Work

A related concept of parameterized testability of graph properties was studied by Iwama and Yoshida [33]. The focus of their work was to design efficient algorithms for the property testing variants of several NP-hard decision problems on graphs, by expressing their complexity in terms of parameters that have been successfully used in the literature on parameterized algorithms. In most of the cases, the parameters that they used are NP-hard to compute. In contrast, our goal is to determine the right input parameters in terms of which to express the complexity of property testers and, more generally, sublinear-time algorithms. The parameters we use are often easy to compute or estimate and, in many situations, can be assumed to be given to the algorithm. We also believe that the parameters that we use are tied to the intrinsic combinatorial structure of the properties and give insights into complexity of testing them.

### 2 The Nonadaptive Sortedness Tester

In this section, we describe a nonadaptive, 1-sided error \(\varepsilon\)-tester for sortedness of arrays containing at most \(r\) distinct values and prove Theorem 1.4. Our tester (Algorithm 1) uses a proximity oblivious tester (POT) for sortedness as a subroutine.

▶ **Definition 2.1 (POT, Goldreich and Ron [29]).** A proximity oblivious tester for a property \(P\) is an algorithm that has oracle access to a function \(f\) and

1. always accepts if \(f \in P\);
2. rejects with probability at least \(\text{dist}(f, P)\) if \(f \notin P\), where \(\text{dist}(f, P)\) is the minimum fraction of values in \(f\) that needs to be changed, so that \(f \in P\).

Observe that a POT for \(P\) can be repeated \(O(1/\varepsilon)\) times to obtain a 1-sided error \(\varepsilon\)-tester for \(P\). We note that Definition 2.1 is a special case of the definition of POT in [29]. Specifically, Goldreich and Ron [29] allow the rejection probability of a POT to be a non-decreasing function of \(\text{dist}(f, P)\). However, the special case in Definition 2.1 is sufficient for our purposes.

We now give an overview of Algorithm 1. It runs for \(O(1/\varepsilon)\) iterations. In each iteration, it first runs a POT for sortedness on a subarray \(B\) of the input array \(A\) consisting of \(1 + 2r/\varepsilon\) (nearly) equally spaced indices. Next, it picks an index \(i \in [n]\) uniformly at random. It compares \(A[i]\) with the array values of the indices closest to \(i\) that were included in \(B\). Algorithm 1 rejects if either of these steps finds elements out of order.

At least three distinct POTs for sortedness of arrays with \(O(\log n)\) query complexity are known [23, 10, 16]. We can use any of them in Algorithm 1. Note that Algorithm 1 is not proximity oblivious itself, as it uses the proximity parameter \(\varepsilon\) to determine its queries. For simplicity, we assume throughout that \(2r/\varepsilon\) is an integer that divides \(n\).

**Proof of Theorem 1.4.** We prove that Algorithm 1 is a nonadaptive, 1-sided error \(\varepsilon\)-tester making \(O(\frac{1}{\varepsilon} \log \frac{n}{\varepsilon})\) queries to test sortedness of arrays with at most \(r\) distinct values. Algorithm 1 is nonadaptive, since its queries can be chosen in advance. It has 1-sided error as it always accepts sorted arrays. Lemmas 2.2 and 2.3 complete the proof of Theorem 1.4.

▶ **Lemma 2.2.** Algorithm 1 makes \(O\left(\frac{1}{\varepsilon} \log \frac{n}{\varepsilon}\right)\) queries.
Algorithm 1: The Nonadaptive Sortedness Tester

input : query access to an array $A$ of size $n$, an upper bound $r$ on the number of distinct values in $A$, and a distance parameter $\varepsilon \in (0,1)$.

1 Let $B$ be the subarray of $A$ consisting of the indices $1, \frac{2n}{2r}, \frac{3n}{2r}, \ldots, \frac{(2r-1)n}{2r}, n$.
   // No need to explicitly construct $B$.
2 repeat $\left\lceil \frac{\varepsilon n}{r} \right\rceil$ times
3 Run a POT for sortedness of arrays (e.g., from [23], [10] or [16]) on $B$ and reject if it rejects.
4 Query an index $i$ from $A$ uniformly at random.
5 Set $k = \left\lfloor \frac{2n}{r} \right\rfloor + 1$. // Note that $B[k] = A\left\lceil \frac{(k-1)n}{2r} \right\rceil$ and $B[k+1] = A\left\lfloor \frac{kn}{2r} \right\rfloor$.
6 Query $B[k]$ and $B[k+1]$.
7 Reject if $(B[k], A[i], B[k+1])$ is not in sorted order.
8 Accept.

Proof. The query complexity of Step 3 is $O(\log |B|) = O(\log(r/\varepsilon))$. Steps 4-7 make a constant number of queries. Steps 3-7 are executed $O(1/\varepsilon)$ times. Hence, the overall query complexity of the tester is $O\left(\frac{1}{\varepsilon} \log \frac{\varepsilon}{r}\right)$.

Recall that an array is $\varepsilon$-far from sorted if at least an $\varepsilon$ fraction of elements need to be modified to make it sorted; otherwise, it is $\varepsilon$-close to sorted.

Lemma 2.3. Algorithm 1, with probability at least 2/3, rejects every array that has at most $r$ distinct values and is $\varepsilon$-far from sorted.

Proof. Consider an array $A$ that has at most $r$ distinct values and is $\varepsilon$-far from sorted. Let $B$ be the subarray of $A$ as defined in Step 1 of Algorithm 1. If $B$ is $\frac{\varepsilon}{r}$-far from sorted, then by the definition of POT for sortedness, Step 3 of our tester rejects with probability at least $\frac{\varepsilon}{r}$ in each iteration. In the rest of the proof, we consider the case when $B$ is $\frac{\varepsilon}{r}$-close to sorted.

Claim 2.4. If $B$ is $\frac{\varepsilon}{r}$-close to sorted, then Steps 4-7 reject with probability at least $\frac{\varepsilon}{r}$ in each iteration.

Proof. The subarray $B$ consists of $1 + 2r/\varepsilon$ (nearly) equally spaced indices, which partition $A$ into $2r/\varepsilon$ intervals of nearly the same size. Let $I = \{I_1, I_2, \ldots, I_{2r/\varepsilon}\}$ denote the set of these intervals. Here, $I_1$ denotes the interval $[2 \cdot \frac{cn}{2r} - 1]$ and, for $k > 1$, the interval $[\frac{(k-1)n}{2r} + 1, \frac{kn}{2r} - 1]$ is denoted by $I_k$. Note that, by definition, $B[k]$ and $B[k+1]$ denote the values of the elements in $A$ present immediately to the left and right of $I_k$, respectively.

An interval $I_k$ is nearly-constant if $B[k] = B[k+1]$. Let $C_t$ be the set of arrays with all their values equal to $t$. Let $A[I_k]$ denote the subarray of $A$ on the indices in $I_k$. Let $d(I_k)$ and $D(I_k)$ denote the fractional and absolute Hamming distance of $A[I_k]$ from the property $C_{B[k]}$. Note that $d(I_k) = D(I_k)/|I_k|$. We now prove Claim 2.4 as follows in two steps. First, we show that $\sum_{I_k \in I'} D(I_k) > \varepsilon n/7$, where $I' = \{I_k \in I : I_k$ is nearly-constant}. Second, we show that Steps 4-7 of Algorithm 1 reject with probability at least $\frac{\varepsilon n}{n}$ in each iteration.

Since $B$ is $\frac{\varepsilon}{r}$-close to sorted, there exists a set $S$ of at most $\varepsilon |B|/7$ indices in $B$ whose values can be changed to make $B$ sorted. Note that, for $r \geq 3$, we have $|S| < r/3$ since

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2 We use $[a..b]$ to denote $\{a, a+1, \ldots, b-1, b\}$ for $a, b \in \mathbb{Z}, a < b$. 

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\(|B| = 1 + 2r/\varepsilon\). Consider the set of intervals \(E_1\) in \(\mathcal{I}\) adjacent to at least one index from \(S\). As each index in \(S\) is adjacent to at most two intervals, \(|E_1| < 2r/3\).

Let \(E_2\) denote the set of intervals in \(\mathcal{I} \setminus E_1\) that are not nearly-constant. For all \(k\) such that \(I_k \in E_2\), we have \(B[k] < B[k+1]\). This is so because, if \(B[k] > B[k+1]\), then \(I_k \in E_1\) and if \(B[k] = B[k+1]\), then \(I_k\) is nearly-constant. The total number of distinct values taken by the elements belonging to intervals in \(E_2\) is at least \(|E_2|\). But \(A\) has at most \(r\) distinct values, and hence, \(|E_2| \leq r\). Consequently, \(|E_1 \cup E_2| < \frac{2r}{3} + r = \frac{5r}{3}\).

Consider the subarray \(A''\) of \(A\) induced by the indices in the intervals in \(\mathcal{I} \setminus (E_1 \cup E_2)\). Let \(D_S(A)\) denote the absolute Hamming distance of the array \(A\) to the sortedness property. As \(D_S(A) \geq \varepsilon n\), we get \(D_S(A'') > \varepsilon n - \frac{5r}{3} \cdot \frac{2r}{3} = \frac{\varepsilon n}{7}\). Note that all the intervals in \(A''\) are nearly-constant. Hence, \((\mathcal{I} \setminus (E_1 \cup E_2)) \subseteq \mathcal{T}'\) and, consequently,

\[\sum_{I_k \in \mathcal{T}'} D(I_k) \geq D_S(A'') > \frac{\varepsilon n}{6} > \frac{\varepsilon n}{7}\]

This completes the first step of the proof.

Consider a nearly-constant interval \(I_k \in \mathcal{T}'\) such that \(D(I_k) > 0\). As \(B[k] = B[k+1]\), there exists \(D(I_k)\) elements in \(I_k\) whose values are not \(B[k]\), i.e.,

\[\{|x \in I_k : A[x] \neq B[k]\}| = D(I_k)\]

Algorithm 1 rejects if it samples an index \(x \in I_k\) in Step 4 such that \(B[k] = B[k+1]\) (i.e., \(I_k \in \mathcal{T}'\)) and \(A[x] \neq B[k]\). As there are \(\sum_{I_k \in \mathcal{T}'} D(I_k)\) such indices in \(A\), Steps 4-7 of Algorithm 1 reject \(A\) with probability at least \(\sum_{I_k \in \mathcal{T}'} D(I_k)/n\). Since \(\sum_{I_k \in \mathcal{T}'} D(I_k) > \varepsilon n/7\), the proof of Claim 2.4 is complete. ▲

Hence, the probability that Steps 3-7 reject in each iteration is at least \(\varepsilon/7\). The probability that Algorithm 1 accepts after \(8\varepsilon/\varepsilon\) iterations is at most \((1-\varepsilon/7)^{8\varepsilon/\varepsilon} \leq e^{-8\varepsilon/7} < 1/3\). This completes the proof of Lemma 2.3. ▲

### 3 The Sortedness Tester with \(O\left(\frac{\log r}{\varepsilon}\right)\) Query Complexity

In this section, we describe a 1-sided error \(\varepsilon\)-tester for sortedness of arrays containing at most \(r\) distinct values and prove Theorem 1.3. The tester, described in Algorithm 2, runs the nonadaptive tester (Algorithm 1) described in Section 2 when \(r \geq 1/\varepsilon\), and a different procedure, which is described in Algorithm 2, otherwise.

**Proof of Theorem 1.3.** We prove that Algorithm 2 is a 1-sided error \(\varepsilon\)-tester making \(O\left(\frac{\log r}{\varepsilon}\right)\) queries to test sortedness of arrays with at most \(r\) distinct values. When \(r \geq 1/\varepsilon\), Algorithm 2 runs Algorithm 1 and outputs its answer. By Theorem 1.4, Algorithm 1 is a 1-sided error \(\varepsilon\)-tester with query complexity \(O\left(\frac{1}{\varepsilon} \log \frac{1}{\varepsilon}\right)\) which is equal to \(O\left(\frac{\log r}{\varepsilon}\right)\) as \(r \geq 1/\varepsilon\). When \(r < 1/\varepsilon\), Algorithm 2 only rejects if it finds array elements out of order, and so, it has 1-sided error. Lemmas 3.1 and 3.2 complete the proof of Theorem 1.3. ▲

**Lemma 3.1.** For \(\varepsilon < 1/\varepsilon\), Algorithm 2 makes \(O\left(\frac{\log r}{\varepsilon}\right)\) queries.

**Proof.** We first bound the query complexity of Steps 4-10. Let \(w\) be the number of times Steps 4-10 are run by Algorithm 2. For \(k \in [0,w]\), let \(T_k\) be the snapshot of the binary search tree \(T\) of array indices (initialized in Step 3) at the end of iteration \(k\). Note that \(T_0 = \{1,n\}\) and \(T_w\) is the tree at the end of the algorithm. Let \(V_k = \{v : v = A[i]\text{ for some } i \in T_k\}\) be the set of all array values of indices in \(T_k\). Observe that once a value \(v\) is in \(V_k\), it
Algorithm 2: The Sortedness Tester

input : query access to an array A of size n, an upper bound r on the number of distinct values in A, and distance parameter ε ∈ (0, 1).
1 If r ≥ 1/ε, run Algorithm 1 and return its answer.
3 Initialize a balanced binary search tree T to contain keys 1 and n.
   // Define successor(i) = min{j ∈ T : j > i}; predecessor(i) = max{j ∈ T : j < i}.
4 while i, j ∈ T such that j = successor(i) and |i − j| > εn/2r and A[i] < A[j] do
   5 Set m = ⌈i + j⌉/2 and query A[m].
   7 If i > 1 and A[predecessor(i)] = A[i] = A[m] then
      Delete i from T.
   8 Delete j from T.
9 Repeat ⌈(2n/3r)⌉ times
   10 Sample an index x from [n] uniformly at random and query A[x].
11 If (A[predecessor(x)], A[x], A[successor(x)]) is not in sorted order then reject.
12 Accept.

remains in V_k for all k’ > k. For v ∈ V_k, define successor-distance(v, T_k) = |i − successor(i)| such that A[i] = v and A[successor(i)] ≠ v, where successor is defined with respect to the tree T_k (for v = A[n], define successor-distance(v, T_k) = 0). Consider the kth iteration of Steps 4-10 where k ∈ [w]. In Step 4 of kth iteration, an index i ∈ T is chosen such that successor-distance(A[i], T_{k-1}) > εn/2r. At the end of the iteration, successor-distance(A[i], T_k) = successor-distance(A[i], T_{k-1})/2 ignoring the errors due to rounding. Generalizing this argument, for each iteration k ∈ [w], there exists some v_k ∈ V_{k-1} \ {A[n]}, such that successor-distance(v_k, T_k) = successor-distance(v_k, T_{k-1})/2.

Fix v* ∈ V_w \ {A[n]}. Let k_1, k_2, . . . , k_q ∈ [w], where k_1 < k_2 < . . . < k_q, be the iterations where the choice of i in Step 4 satisfies A[i] = v*. From the description of the tester, for any i ∈ [2,q], we have successor-distance(v*, T_{k_i}) = successor-distance(v*, T_{k_{i-1}})/2. By extending this relation, we get successor-distance(v*, T_{k_q}) = successor-distance(v*, T_{k_{q-1}})/2^{q-1}. But successor-distance(v*, T_{k_q}) < n and εn/4r < successor-distance(v*, T_{k_q}) ≤ εn/2r. Solving for q, we get

\[ 2^{q-1} \frac{\text{successor-distance}(v^*, T_{k_q})}{\text{successor-distance}(v^*, T_{k_{q-1}})} < \frac{n}{εn/4r} = \frac{4r}{ε}; \]

\[ q < \log \frac{8r}{ε}. \]

Hence, the tester runs at most log(8r/ε) iterations where successor-distance(v*, ·) is halved. Accounting for all the iterations for each value in V_w \ {A[n]}, we get

\[ w < |V_w| \cdot \log(8r/ε) \leq r \log(8r/ε), \]

since |V_w| ≤ r. In each iteration, the tester makes a constant number of queries. So, the overall query complexity of Steps 4-10 is O(r log 2/ε). The query complexity of Steps 11-13 is O(1/ε). Hence, the overall query complexity of the tester is O \left( \frac{1}{ε} + r \log \frac{2}{ε} \right).

Now, we prove that O \left( r \log \frac{2}{ε} \right) = O \left( \frac{\log \frac{2}{ε}}{r} \right) for r < 1/ε. We have O \left( r \log \frac{2}{ε} \right) = O \left( r \log \frac{1}{ε} \right) as r < 1/ε. Note that the function g(x) = \frac{1}{\log x} is increasing for x ≥ 3. Hence, for r < 1/ε,
we have $\frac{r}{\log r} < \frac{1/\varepsilon}{\log(1/\varepsilon)}$, and hence $r \log \frac{1}{\varepsilon} < \frac{\log r}{\varepsilon}$. Therefore, the query complexity of Algorithm 2 is $O\left(\frac{\log r}{\varepsilon}\right)$.

Lemma 3.2. Steps 2-14 of Algorithm 2, with probability at least $2/3$, reject every array that has at most $r$ distinct values and is $\varepsilon$-far from sorted, when $r < 1/\varepsilon$.

Proof. Consider an array $A$ that has at most $r$ distinct values and is $\varepsilon$-far from sorted, where $r < 1/\varepsilon$. Algorithm 2 rejects whenever it finds elements out of order. We show that Steps 11-13 reject with probability at least $2/3$, if Steps 2-10 do not find array elements out of order.

Consider the indices in the tree $T$ at the end of the while loop. Let $E = \{j \in T : A[j] < A[\text{successor}(j)]\}$ be the indices in $T$ whose array values differ from that of their respective successor in $T$. As $A$ has at most $r$ distinct values, by Pigeonhole principle, $|E| < r$. Each $i \in E$ satisfies $|i - \text{successor}(i)| \leq \varepsilon n/2r$. Define $E' = \{k \in [n] : i < k < \text{successor}(i), i \in E\}$. Clearly, $|E'| \leq \frac{c \varepsilon}{2r} |E| < \frac{c \varepsilon}{2r} n$. Consider the subarray of $A$ indexed by $[n] \setminus E'$. This subarray is $\frac{2}{3}$-far from sorted as $A$ is $\varepsilon$-far from sorted. Also, all $k \in [n] \setminus E'$ satisfy $\text{predecessor}(k) < k < \text{successor}(k)$ and $A[\text{predecessor}(k)] = A[\text{successor}(k)]$ (note that the definitions of predecessor and successor are applicable to all elements in $[n]$). That is, for all such indices $k$, we know what the element $A[k]$ should be if $A$ is sorted. Recall that if $A[i] = A[j]$, then $[i, j]$ constitutes a nearly-constant interval, as defined in Section 2. By the proof method used in Lemma 2.3, there exists at least $\varepsilon n/2$ indices of the form $k \in [n]$ such that $A[\text{predecessor}(k)] = A[\text{successor}(k)]$ and $A[k] \neq A[\text{successor}(k)]$. The probability that Steps 12 and 13 fail to capture such an index in any of its $\left\lceil \frac{2 \ln 3}{\varepsilon} \right\rceil$ iterations is at most

$$
(1 - \varepsilon/2)^{\frac{2 \ln 3}{\varepsilon}} \leq 1/3.
$$

4 The Monotonicity Tester over Hypergrids

In this section, we describe a monotonicity tester for functions over hypergrid domains and prove Theorem 1.7. We prove the correctness of this tester using the correctness of the sortedness tester described in Section 3, a dimension reduction theorem by Chakrabarty et al. [15] and the work investment strategy by Berman et al. [9].

An axis-parallel line $\ell$ of the hypergrid $[n]^d$ is a set of $n$ points that agree on all but one coordinate. Let $f|_{\ell}$ denote the restriction of a function $f$ to $\ell$. Note that $f|_{\ell}$ can be thought of as a real-valued function over $[n]$.

Algorithm 3: The Monotonicity Tester over Hypergrids

```plaintext
input : query access to $f : [n]^d \mapsto \mathbb{R}$, an upper bound $r$ on $|\text{Im}(f)|$, and a distance parameter $\varepsilon \in (0, 1).

1 for $i = 1$ to $\left\lceil 3 \log \frac{4d}{\varepsilon} \right\rceil$ do
2   repeat $\left\lceil \frac{4d \ln 4}{d \varepsilon} \right\rceil$ times
3     Sample a uniformly random axis-parallel line $\ell$.
4     Repeat twice: run Algorithm 2 on the array induced by $f|_{\ell}$, with the distance parameter set to $2^{-i}$ and the upper bound on the number of distinct elements set to $r$; reject if it rejects at least once.
5 Accept.
```
The tester iteratively samples uniformly random axis-parallel lines, runs Algorithm 2 on each of them, and rejects if any run of Algorithm 2 rejects. We now analyze the tester and prove Theorem 1.7.

**Proof of Theorem 1.7.** We prove that Algorithm 3 is a 1-sided error $\varepsilon$-tester that makes $O\left(\frac{\varepsilon}{\varepsilon^2} \log \frac{\varepsilon}{\varepsilon} \log r\right)$ queries to test monotonicity of real-valued functions $f : [n]^d \to \mathbb{R}$ over the hypergrid domain, where $|\text{Im}(f)| \leq r$. Algorithm 3 has 1-sided error because Algorithm 2, which it runs as a subroutine, has 1-sided error. Lemmas 4.1 and 4.2 complete the proof of Theorem 1.7.

**Lemma 4.1.** Algorithm 3 makes $O\left(\frac{\varepsilon}{\varepsilon^2} \log \frac{\varepsilon}{\varepsilon} \log r\right)$ queries.

**Proof.** The query complexity of a single execution of Step 4 during the $i$th iteration of the outermost loop (Step 1) is $O(2^i \log r)$. As Step 4 is repeated $O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\varepsilon}\right)$ times in the $i$th iteration, the overall query complexity of the $i$th iteration of the tester is $O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\varepsilon} \log \frac{\varepsilon}{\varepsilon} \log r\right)$. The outermost loop is executed $O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\varepsilon}\right)$ times, and hence the query complexity of the tester is $O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\varepsilon} \log \frac{\varepsilon}{\varepsilon} \log r\right)$.

**Lemma 4.2.** Algorithm 3, with probability at least 2/3, rejects every function over the hypergrid domain which is $\varepsilon$-far from sorted and has image size is at most $r$.

**Proof.** Let $f : [n]^d \to \mathbb{R}$ be $\varepsilon$-far from monotone, with $|\text{Im}(f)| \leq r$. Let $\mathcal{L}_{n,d}$ denote the set of all axis-parallel lines in $[n]^d$ and $d_M(f)$ denote the relative distance of $f$ to monotonicity. We also use $d_M(f|\ell)$ to denote the relative distance to monotonicity of the function $f|\ell$. We have $|\text{Im}(f|\ell)| \leq r$ since $|\text{Im}(f)| \leq r$. We use the following dimension reduction theorem proved by Chakrabarty et al. [15].

**Theorem 4.3** (Chakrabarty et al. [15]).

$$
\mathbb{E}_{\ell \sim \mathcal{L}_{n,d}}[d_M(f|\ell)] \geq \frac{d_M(f)}{4d}.
$$

We note that Theorem 4.3 is a special case of the dimension reduction theorem proved in [15]. Clearly, if $d_M(f) \geq \varepsilon$, then, $\mathbb{E}_{\ell \sim \mathcal{L}_{n,d}}[d_M(f|\ell)] \geq \varepsilon/4d$. We use the work investment strategy due to Berman et al. [9] to extend the monotonicity tester on the line domain to the hypergrid domain.

**Theorem 4.4** (Berman et al. [9]). For a random variable $X \in [0,1]$ with $\mathbb{E}[X] \geq \mu$ for $\mu < 1/2$, let $p_i = \Pr[X \geq \frac{1}{i}]$ and $\delta \in (0,1)$ be the desired probability of error. Let $k_i = \frac{4 \ln 1/\delta}{2^i \mu}$. Then,

$$
\prod_{i=1}^{\left\lceil 3 \log (1/\mu) \right\rceil} (1 - p_i)^{k_i} \leq \delta.
$$

Consider running Algorithm 3 on $f$. Let $X = d_M(f|\ell)$, where $\ell$ is sampled uniformly at random from $\mathcal{L}_{n,d}$. We apply the work investment strategy (Theorem 4.4) on $X$ with error probability $\delta = 1/4$. By Theorem 4.3, $\mathbb{E}[X] \geq \varepsilon/4d$. Thus, in Theorem 4.4, we set $\mu = \varepsilon/4d$ and $k_i = \frac{4 \ln \delta}{2^i \mu}$ for all $i \in \left\lceil 3 \log \frac{4d}{\varepsilon} \right\rceil$. By Theorem 4.4, with probability at least 3/4, for some $i \in \left\lceil 3 \log \frac{4d}{\varepsilon} \right\rceil$, we sample a line $\ell$ such that $d_M(f|\ell) \geq 2^{-i}$ in Step 3. Conditioned on sampling such a line, Step 4 rejects $\ell$ with probability at least $8/9$. Thus, given a function $f$ that is $\varepsilon$-far from sorted, Algorithm 3 rejects $f$ with probability at least $\frac{3}{4} \cdot \frac{8}{9} = \frac{2}{3}$, as required. This completes the proof of Lemma 4.2. ◁
Note on a nonadaptive tester for hypergrids. We can get a nonadaptive, 1-sided error \( \varepsilon \)-tester for monotonicity over hypergrids by using Algorithm 1 instead of Algorithm 2 in Step 4 of Algorithm 3. The same analysis goes through for this case and the overall query complexity of the tester is \( O \left( \frac{d}{\varepsilon} \log \frac{d}{\varepsilon} \log \frac{rd}{\varepsilon} \right) \).

5 The Uniform Tester for Sortedness

In this section, we describe a nonadaptive \( \varepsilon \)-tester that makes \( O(\sqrt{r}/\varepsilon) \) uniform and independent queries to test sortedness of arrays containing at most \( r \) distinct values and prove Theorem 1.5.

Our tester is Algorithm 4. The bound on the query complexity of the tester follows directly from its description. The tester has 1-sided error as it always accepts sorted arrays. In the rest of the section, we show that, with high probability, the tester rejects arrays that are \( \varepsilon \)-far from sorted.

Algorithm 4: The Uniform Sortedness Tester

**input** : query access to an array \( A \) of size \( n \), an upper bound \( r \) on the number of distinct values in \( A \), and a distance parameter \( \varepsilon \in (0,1) \).

1. Sample \( \left\lceil \frac{24\sqrt{r}}{\varepsilon^2} \right\rceil \) indices from \( A \) uniformly and independently at random.
2. **Reject** if the array restricted to the sampled indices is not sorted; otherwise, **accept**.

**Lemma 5.1.** Algorithm 4, with probability at least \( 2/3 \), rejects every array that has at most \( r \) distinct elements and is \( \varepsilon \)-far from sorted.

**Proof.** Consider an array \( A \) that has at most \( r \) distinct values and is \( \varepsilon \)-far from sorted. Recall that a pair of indices \( (x, y) \), where \( x, y \in [n] \) and \( x < y \), is violated in an array \( A \) if \( A(x) > A(y) \). Consider the undirected violation graph \( G = ([n], E) \) of \( A \), where an edge \( (u, v) \in E \) if \( (u, v) \) is a violated pair. Dodis et al. [22, Lemma 7] show that if \( A \) is \( \varepsilon \)-far from sorted then \( G \) has a matching \( M \) of size at least \( \varepsilon n/2 \).

For a pair \( (x, y) \in [n] \times [n] \) such that \( x < y \), we refer to \( x \) as its lower endpoint and \( y \) as its higher endpoint. We first partition the pairs in \( M \) into \( r \) classes as follows. Let \( v_1 < v_2 < \cdots < v_r \) be the values in the range. A pair \( (x, y) \in M \) such that \( x < y \) belongs to the \( i \)th class \( C_i \), if \( A(x) = v_i \). Note that \( C_1 \) is empty. For each \( i \in [r] \), let \( C_l^i \) and \( C_r^i \) denote the set of lower and higher endpoints of pairs in \( C_i \), respectively. Note that \( |C_i| = |C_l^i| = |C_r^i| \). For each \( i \in [r] \), define the \( i \)th lower bucket \( B_l^i \) to consist of the smallest \( \lceil |C_l^i|/2 \rceil \) indices in \( C_l^i \) and the \( i \)th higher bucket \( B_r^i \) to consist of the largest \( \lceil |C_l^i|/2 \rceil \) indices in \( C_r^i \). Note that \( \bigcup_{i\in[r]} B_l^i = \bigcup_{i\in[r]} B_r^i \geq \varepsilon n/4 \). It is easy to see that for each \( i \in [r] \), every pair in \( B_l^i \times B_r^i \) is a violated pair. Therefore, if an algorithm samples indices from both \( B_l^i \) and \( B_r^i \), for some \( i \in [r] \), it rejects. To bound the probability of this event from below, we use the following generalization of the Birthday Paradox proved by Goldreich et al. [27, Lemma 19].

**Claim 5.2 ([27, Lemma 19]).** Let \( S_1, S_2, \ldots, S_r, T_1, T_2, \ldots, T_r \) be disjoint subsets of a universe \( U \). For each \( i \in [r] \), let \( p_i = |S_i|/|U| \) and \( q_i = |T_i|/|U| \). Let \( \rho = \sum_i \min\{p_i, q_i\} \). Then, if we uniformly sample \( 6\sqrt{\rho}/\rho \) elements from \( U \), with probability at least 2/3, for some \( i \in [r] \), the sample will contain at least one element from both \( S_i \) and \( T_i \).
If we set $S_i = B_i^L$ and $T_i = B_i^H$ for each $i \in [r]$ in Claim 5.2, we have $\rho \geq \varepsilon / 4$. Therefore, a uniform sample of $2^{\sqrt{r}} / \varepsilon$ points from $[n]$, with probability at least $2/3$, will have, for some $i \in [r]$, an index from $B_i^L$ and $B_i^H$, and the algorithm will reject. This completes the proof of the lemma.

6 A Lower Bound for the Uniform Sortedness Tester

In this section, we prove that $\Omega(\sqrt{r})$ uniform queries are required to test sortedness of an array with at most $r$ distinct values, even when one allows for 2-sided error, and prove Theorem 1.6. The proof uses Yao’s principle [41], the version with two distributions (see, e.g., Raskhodnikova and Smith [39]). We first define two hard distributions $\mathcal{P}$ and $\mathcal{N}$ on arrays with $r$ distinct values such that every array drawn from $\mathcal{P}$ is in sorted order and every array drawn from $\mathcal{N}$ is $\frac{1}{\varepsilon}$-far from sorted. We then show that, for any tester that uses $o(\sqrt{r})$ uniform queries, the statistical difference between tester’s views of the two distributions is small, and hence, with high probability, it cannot distinguish between the distributions.

The statistical distance between two distributions $\mathcal{D}_1$ and $\mathcal{D}_2$, denoted by $SD(\mathcal{D}_1, \mathcal{D}_2)$, is defined as

$$SD(\mathcal{D}_1, \mathcal{D}_2) = \max_{S \subseteq (\text{support}(\mathcal{D}_1): \text{support}(\mathcal{D}_2))} \left| \Pr_{x \sim \mathcal{D}_1} [x \in S] - \Pr_{x \sim \mathcal{D}_2} [x \in S] \right|.$$

We write $\mathcal{D}_1 \approx \delta \mathcal{D}_2$ to denote $SD(\mathcal{D}_1, \mathcal{D}_2) \leq \delta$.

**Proof of Theorem 1.6.** First, we define distributions $\mathcal{P}$ and $\mathcal{N}$ on arrays of size $n$ taking values in the set $[r]$, where $n \geq 16r \ln 8r$. Without loss of generality, we assume that $r$ is an even number that divides $n$.

The distribution $\mathcal{P}$ is constructed as follows. Partition an $n$-element array into $r/2$ blocks, each of length $2n/r$. For $i \in [r/2]$, set the values in the $i^{th}$ block of the array to $(2i - 1), 2i, 2i, \ldots, 2i$ with probability $1/2$ and $(2i - 1), (2i - 1), \ldots, (2i - 1), 2i$ with probability $1/2$, independent of the other blocks.

The distribution $\mathcal{N}$ is constructed as follows. As before, partition an $n$-element array into $r/2$ blocks, each of length $2n/r$. For $i \in [r/2]$, set the first value in the $i^{th}$ block to $(2i - 1)$ and the last value to $2i$. The values at all other indices in that block are set to either $(2i - 1)$ or $2i$ uniformly and independently at random.

Note that every array drawn from $\mathcal{P}$ is in sorted order. We will show that, with high probability, an array drawn from $\mathcal{N}$ is $\frac{1}{\varepsilon}$-far from sorted.

**Lemma 6.1.** Let $E$ denote the event that an array chosen according to $\mathcal{N}$ is $\frac{1}{\varepsilon}$-far from sorted. Then,

$$Pr[E] > \frac{5}{6}.$$

**Proof.** Consider an array $A$ chosen according to $\mathcal{N}$. Consider the $i^{th}$ block of $A$ for some $i \in [r/2]$. Let $Y_{2i}$ denote the number of elements with value $2i$ in the first half of this block and $Y_{2i-1}$ denote the number of elements with value $(2i - 1)$ in the second half of the block. As the size of each half of the block is $n/r$, and the value at each index (except for the first and the last index) is assigned either $(2i - 1)$ or $2i$ uniformly and independently at random,

$$\mathbb{E}[Y_{2i}] = \mathbb{E}[Y_{2i-1}] = \frac{n}{2r} - \frac{1}{2}.$$
By a Chernoff bound, for all \( i \in \left[ \frac{r}{2} \right] \),
\[
\Pr \left[ Y_{2i} \leq \frac{n}{4r} \right] = \Pr \left[ Y_{2i-1} \leq \frac{n}{4r} \right] = \Pr \left[ Y_{2i-1} \leq \left( 1 - \frac{n - 2r}{2(n - r)} \right) \left( \frac{n}{2r} - \frac{1}{2} \right) \right] \\
\leq \exp \left( -\frac{n}{16r} \cdot \frac{(n - 2r)^2}{n(n - r)} \right) \\
< \frac{1}{6r}.
\]

If \( Y_{2i} > n/4r \) and \( Y_{2i-1} > n/4r \), then at least \( n/4r \) elements in \( r \)-th block need to be changed to make it sorted, as all the indices with value \( 2i \) in the first half or all the indices with value \( 2i - 1 \) in the last half need to be changed. By the union bound,
\[
\Pr \left[ \bigvee_{j=1}^{r} \left( Y_j \leq \frac{n}{4r} \right) \right] \leq r \cdot \Pr \left[ Y_1 \leq \frac{n}{4r} \right] < \frac{1}{6}.
\]

With probability at least 5/6, we have \( Y_{2i} > n/4r \) and \( Y_{2i-1} > n/4r \) for all \( i \in [r/2] \). This implies that at least \( n/4r \) elements need to be changed in each of the \( r/2 \) blocks to make it sorted. Hence, with probability at least 5/6, the array \( A \) is \( \frac{1}{5} \)-far from sorted. ▶

Denote the conditional distribution \( \mathcal{N}|_E \) by \( \hat{\mathcal{N}} \). Any instance sampled according to \( \hat{\mathcal{N}} \) is \( \frac{1}{5} \)-far from sorted. The statistical distance \( \text{SD}(\mathcal{N}, \hat{\mathcal{N}}) \) can be bounded using the following lemma proven by Raskhodnikova and Smith [39].

▶ **Lemma 6.2** ([39, Claim 4]). Let \( E \) be an event that happens with probability at least \( 1 - \delta \) under the distribution \( \mathcal{D} \). Then, \( \mathcal{D} \approx_{\delta'} \mathcal{D}|_E \), where \( \delta' = \frac{1}{1-\delta} - 1 \).

Applying Lemma 6.2 to \( \mathcal{N} \) and \( \hat{\mathcal{N}} \), we get \( \mathcal{N} \approx_{1/5} \hat{\mathcal{N}} \).

Consider any \( \frac{1}{5} \)-tester for sortedness that makes \( q \) queries where \( q \leq \sqrt{q}/5 \). Define \( \mathcal{P} \)-view to be the distribution of values at the \( q \) locations queried by the tester in an array sampled according to \( \mathcal{P} \). Similarly, define \( \mathcal{N} \)-view and \( \hat{\mathcal{N}} \)-view. Next, we show that it is hard to distinguish \( \mathcal{P} \)-view from \( \hat{\mathcal{N}} \)-view.

▶ **Lemma 6.3.**
\[
\text{SD}(\mathcal{P} \text{-view}, \hat{\mathcal{N}} \text{-view}) < \frac{1}{3}.
\]

**Proof.** Let \( F \) denote the event that at least 2 out of the tester’s \( q \) uniform samples from an array \( A \) are from the same block. An upper bound on the probability of this event can be obtained using the following lemma.

▶ **Lemma 6.4** (Bellare and Rogaway [3]). Consider \( q \) balls and \( N \) bins, where each ball is assigned uniformly and independently at random to one of the bins. The probability that there exists a pair of balls assigned to the same bin is at most \( \frac{2(q-1)}{2N} \).

By Lemma 6.4, we get \( \Pr[F] \leq \frac{q(q-1)}{2q^2} < \frac{q}{q} = \frac{1}{25} \). Then, by Lemma 6.2,
\[
\mathcal{P} \text{-view} \approx_{1/24} \mathcal{P} \text{-view}|_F; \quad \text{\quad (1)}
\]
\[
\mathcal{N} \text{-view} \approx_{1/24} \mathcal{N} \text{-view}|_F. \quad \text{\quad (2)}
\]

Since \( \mathcal{N} \approx_{1/5} \hat{\mathcal{N}} \), the definition of statistical difference implies that
\[
\mathcal{N} \text{-view} \approx_{1/5} \hat{\mathcal{N}} \text{-view}. \quad \text{\quad (3)}
\]
It remains to show that $\mathcal{P}\text{-view}\mid_{\mathcal{P}} = \mathcal{N}\text{-view}\mid_{\mathcal{P}}$. Let $x$ be an index in the $i^{\text{th}}$ block, for some $i \in [r/2]$. If $x$ is neither the first nor the last index of $i^{\text{th}}$ block, then $\Pr[A[x] = (2i − 1)] = \Pr[A[x] = 2i] = 1/2$ irrespective of whether $A \leftarrow \mathcal{P}$ or $A \leftarrow \mathcal{N}$. If $x$ is the first or the last index of the $i^{\text{th}}$ block, then $A[x]$ is fixed to the same value under both $\mathcal{P}$ and $\mathcal{N}$. If $\mathcal{P}$ holds, then at most 1 index from each block is sampled by the tester. By the definition of $\mathcal{P}$ and $\mathcal{N}$, for any two indices from different blocks, the values assigned to them are independent of each other. Hence, $\mathcal{P}\text{-view}\mid_{\mathcal{P}} = \mathcal{N}\text{-view}\mid_{\mathcal{P}}$. By (1)-(3),

$$\text{SD}(\mathcal{P}\text{-view}, \hat{\mathcal{N}}\text{-view}) \leq \frac{1}{24} + \frac{1}{24} + \frac{1}{5} < \frac{1}{3}.$$  

This completes the proof of Lemma 6.3. \hfill $\blacksquare$

By Yao’s principle [41], as stated in [39, Claim 5], for $q \leq \sqrt{r}/5$, it is hard for any $\frac{1}{q}$-tester using $q$ uniform queries to distinguish $\mathcal{P}$ from $\hat{\mathcal{N}}$. Thus, uniform testers for sortedness of arrays with values in $[r]$ require $\Omega(\sqrt{r})$ queries. This completes the proof of Theorem 1.6. \hfill $\blacksquare$

### 7 Testing Convexity

In this section, we describe a nonadaptive tester for convexity of functions $f : [n] \mapsto \mathbb{R}$ and prove Theorem 1.8. Recall that a function $f : [n] \mapsto \mathbb{R}$ is convex if $f(i) − f(i−1) \leq f(i+1) − f(i)$ for $1 \leq i \leq n$. Our convexity tester is Algorithm 5. It uses the nonadaptive convexity tester of Parnas et al. [37] as a black box.

**Algorithm 5:** The Convexity Tester

```plaintext
input : query access to $f : [n] \mapsto \mathbb{R}$, an upper bound $r$ on $|\text{Im}(f)|$, and a distance parameter $\varepsilon \in (0, 1)$.

if $r \geq \frac{\varepsilon n}{4}$ then
1       Run the $\varepsilon$-tester for convexity by Parnas et al. [37] on $f$ and reject if it rejects.
else
2       Let $M \leftarrow [r+1, \ldots, n-r]$.
3       Sample $\lceil \frac{r}{5} \rceil$ indices from $M$ uniformly and independently at random.
4       Reject if $f$ restricted to those indices is not constant.
5       Accept.
```

The query complexity of our tester is $O(1/\varepsilon)$ when $r < \varepsilon n/3$, as is evident from its description. In the other case, $n \leq 3r/\varepsilon$, our tester runs the tester of [37], which makes $O(\log n/\varepsilon)$ queries. Substituting the upper bound on $n$, we get the query complexity bound claimed in Theorem 1.8.

Given a function $f : [n] \mapsto \mathbb{R}$ and a set $S \subseteq [n]$, let $f_S$ denote the restriction of $f$ to the indices in $S$ whenever $S \neq \emptyset$. To prove the correctness of our tester, we first prove the following characterization of convex functions with image size at most $r$.

**Claim 7.1.** If $f : [n] \mapsto \mathbb{R}$ is convex and $|\text{Im}(f)| \leq r$, then $f_M$ for $M = [r+1..n-r]$ is a constant function.

**Proof.** We can assume that $r < n/2$, for otherwise, $M = \emptyset$. Assume for the sake of contradiction that there exists points $x, x+1 \in M$ such that $f_M(x) \neq f_M(x + 1)$. If $f_M(x) < f_M(x + 1)$, then $f$ has to be monotonically increasing on the domain restricted to $[x+1, \ldots, n]$, which has more than $r$ elements in it as $x < n-r+1$. By the pigeonhole
principle, this results in a contradiction, as $|\text{Im}(f)| \leq r$. If $f_M(x) > f_M(x+1)$, then $f$ has to be monotonically decreasing on the set $[1, \ldots, x+1]$, which has more than $r$ elements in it since $x \geq r$. By the pigeonhole principle, this also leads to a contradiction, as $|\text{Im}(f)| \leq r$. Hence, $f$ can take only one value on $M$ and therefore, $f_M$ is a constant function.

We will now show that the tester accepts every function that is convex and rejects with probability at least $2/3$, every function that is $\epsilon$-far from convex.

Lemma 7.2. Consider a function $f : [n] \mapsto \mathbb{R}$. Algorithm 5, on input $r \geq |\text{Im}(f)|$ and $\epsilon$, accepts if $f$ is convex and rejects, with probability at least $2/3$, if $f$ is $\epsilon$-far from convex.

Proof. If $r \geq \frac{3n}{2}$, Algorithm 5 runs the tester for convexity by [37], and so the correctness follows from their analysis.

Consider the case where $r < \frac{3n}{2}$. It follows from Claim 7.1 that Algorithm 5 accepts $f$ if it is convex. Now assume that $f$ is $\epsilon$-far from convex. It remains to prove that $f_M$ is $\epsilon/3$-far from being a constant function, where $M = [r+1, \ldots, n-r]$. Assume for the sake of contradiction that $f_M$ is $\epsilon/3$-close to constant. We will construct a convex function $g : [n] \mapsto \mathbb{R}$ such that $g$ is $\epsilon$-close to $f$ and satisfies $|\text{Im}(g)| \leq r$. This will give us the required contradiction. Let the constant function closest to $f_M$ be $h$, where $h(x) = c$ for every $x \in M$. The function $g$ is then defined as a constant function taking the value $c$ on all points in $[n]$. Since the Hamming distance of $f_M$ from $h$ is at most $\epsilon n/3$, the total Hamming distance of $f$ from $g$ is at most $\epsilon n/3 + 2r < \epsilon n$. This contradicts the fact that $f$ is $\epsilon$-far from convex. Hence, $f_M$ is $\epsilon/3$-far from being a constant function. The probability that $4/\epsilon$ samples fail to detect that $f_M$ is $\epsilon/3$-far from constant is at most $(1-\epsilon/3)^{4/\epsilon} \leq \exp(-4/3) < 1/3$. ▷

References

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